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## Stock Assessment for the Southern and Eastern Scalefish and Shark Fishery: 2020 and 2021


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## Cover photographs

Front cover, jackass morwong, orange roughy, blue grenadier, and flathead.

## Report structure

Part 1 of this report describes the Tier 1 assessments of 2021. Part 2 describes the Tier 4 and Tier 5 assessments, catch rate standardisations and other work contributing to the assessment and management of SESSF stocks in 2021.

# Stock Assessment for the Southern and Eastern Scalefish and Shark Fishery 2020 and 2021 

Part 1: 2021
G.N. Tuck

May 2022
Report 2019/0800
Australian Fisheries Management Authority

## Stock Assessment for the Southern and Eastern Scalefish and Shark Fishery: 2021

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## 6. Blue Grenadier (Macruronus novaezelandiae) stock assessment based on data up to 2020

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### 6.1 Executive Summary

This document presents the agreed base case for an updated quantitative Tier 1 assessment of Blue Grenadier (Macruronus novaezelandiae) for presentation at the SERAG3 meeting in 2021. The last full assessment was conducted in 2018 (Castillo-Jordán and Tuck, 2018b). The preliminary base case was presented at SERAG2 (October 2021; Tuck and Bessell-Browne, 2021) and the 2018 assessment was updated by the inclusion of data up to the end of 2020, which entails an additional three years of catch, discard, CPUE, length and age data and ageing error updates. The development of, and results from, the preliminary base case for Blue Grenadier through the sequential updating of recent data in the stock assessment, using the stock assessment package Stock Synthesis (SS-V3.30, Methot and Wetzel (2013)) is described in Tuck and Bessell-Browne (2021) and is not repeated here. This document describes the agreed base case from SERAG2 which differs from the preliminary base case through the inclusion of estimation of both female and male natural mortality, and no longer including the FIS survey results.

Results of the base case show reasonably good fits to the length-composition data, conditional age at length, egg and acoustic surveys and discard mass. As has been noted in previous Blue Grenadier assessments, the fit to the standardized non-spawning catch-rate index is generally poor; the model is unable to fit to the high early catch rates and over-estimates catch rates during the early 2000s. More recent catch rates fit reasonably well, including the recent marked increase in catch rate in 2019 and 2020.

The estimated time series of recruitment under the base-case parameter set shows the typical episodic nature of Blue Grenadier recruitment, with strong year-classes in 1979, the mid-1980s, 1994, and 2003, with very little recruitment between these years. However, recent recruitments are more stable, as was first observed in the 2018 assessment. The trajectories of spawning biomass show increases and decreases in spawning biomass as strong cohorts move into and out of the spawning population.

For the base case model, the estimated virgin female spawning biomass ( $S S B_{0}$ ) is 37,445 tonnes and the projected 2022 spawning stock biomass will be $155 \%$ of $S S B_{0}$ (projected assuming 2020 catches in 2021), compared to $122 \%$ for 2019 in the 2018 assessment. The 2022 recommended biological catch (RBC) under the 20:35:48 harvest control rule is $23,777 \mathrm{t}$, with 245 t estimated discards ( $23,532 \mathrm{t}$ retained). The long-term RBC is $7,100 \mathrm{t}$, with 183 t discards.

### 6.2 Introduction

An integrated analysis model, implemented in the generalized stock assessment software package, Stock Synthesis (SS) (Methot and Wetzel, 2013), was applied to the stock of Blue Grenadier in the Southern and Eastern Scalefish and Shark Fishery (SESSF), with data updated by the inclusion of data
up to the end of the 2020 calendar year (length-composition and conditional age-at-length data; age reading-error matrices, standardized catch rate series; landings and discard catch weight) and information from acoustic surveys of spawning biomass (series from 2003-2010, pertaining to total spawning biomass), with an assumption of 2-times turnover on the spawning ground (Russell and Smith, 2006; Punt et al., 2015). The base-case egg survey estimates of female (only) spawning biomass for 1994 and 1995 are included. The model fits directly to length-composition data (by sex where possible) and conditional age-at-length data by fleet. Retained length-composition data from port and onboard samples are fit separately with a common selectivity curve by fleet.

The assessment model presented in 2011 (Tuck and Whitten, 2011; Tuck, 2011) was the first for Blue Grenadier to be implemented using Stock Synthesis (SS). The 2013 assessment updated this assessment using SS-V3.22a (Tuck, 2013), and the last full assessment was in 2018 (Castillo-Jordán and Tuck, 2018b), using 3.30.12.00-safe. The preliminary base case presented to SERAG in October 2021 (Tuck and Bessell-Browne, 2021) illustrated the changes that have occurred since 2018 through changes to software, assessment practices and new data (bridging). The bridging analysis are not repeated here.

The use of SS allows for multiple fishing fleets and can fit simultaneously to several data sources and types of information. The population dynamics model, and the statistical approach used in the fitting of the model to the various types of data, is outlined fully in the SS user manual (Methot et al., 2021) and is not reproduced here. This document updates the assessment presented in 2018 and the preliminary assessment presented at SERAG in October 2021 (Tuck and Bessell-Browne, 2021).

### 6.3 The fishery

Blue Grenadier are found from New South Wales around southern Australia to Western Australia, including the coast of Tasmania. Blue Grenadier is a moderately long-lived species with a maximum age of about 25 years. Age at maturity is approximately four years for males and five years for females (length-at- $50 \%$ maturity for females is 57 cm and 64 cm respectively) based upon 32,000 Blue Grenadier sampled between February 1999 and October 2001 (Russell and Smith, 2006). There is also evidence that availability to the gear on the spawning ground differs by sex, with a higher proportion of small males being caught than females. This is most likely due to the arrival of males on the spawning ground at a smaller size (and younger age) than females. This was also noted by Russell and Smith (2006) who state that "young males entered the fishery one year earlier than females" and is consistent with information for Hoki from New Zealand (Annala et al., 2003). Large fish arrive earlier in the spawning season than small fish. Spawning occurs predominantly off western Tasmania in winter (the peak spawning period based upon mean gonadosomatic index (GSI) calculated by month was estimated to be between June and August according to Russell and Smith (2006)). There is some evidence that a high proportion of fish remain spawning in September. Variations in spawning period noted by Gunn et al. (1989) may occur due to inter-annual differences in the development of coastal current patterns around Tasmania. Adults disperse following the spawning season and while fish are found throughout the south east region during the non-spawning season, their range is not well defined. Spawning fish have been caught off the east coast of Australia, and larvae from a likely eastern spawning area have been described by Bruce et al. (2001). Blue Grenadier are caught by demersal trawling. There are two defined fleets: the spawning (SESSF Zone 40, months June, July and August) and non-spawning fisheries (all other months and zones).

### 6.4 Data

The assessment has been updated since the previous assessment (Castillo-Jordán and Tuck, 2018) by including recent length-composition and conditional age-at-length data from the spawning and nonspawning fisheries; updated standardized CPUE series (Sporcic, 2021), the total mass landed and discarded, and updated age-reading error matrices. Acoustic estimates of spawning biomass (20032010) and estimates of the female spawning biomass in 1994 and 1995 from egg surveys (Bulman et al., 1999) are included (Figure 6.1). The agreed base case no longer includes the FIS abundance estimates from the non-spawning area, as SERAG2 did not believe the series (FIS1-3) was indexing either the spawning or non-spawning biomass; extremely large inter- annual fluctuations in survey biomass are evident. Data were formulated by calendar year (i.e. 1 Jan to 31 Dec ), as in previous models.

Data by type and year, circle area is relative to precision within data type


Figure 6.1. A summary of the input data for the base case Blue Grenadier assessment.

### 6.4.1 Catch data

### 6.4.1.1 Landings

The landings from the logbook data were used to apportion catches to the spawning and non-spawning fisheries (Table 6.1). The logbook landings have been adjusted upwards to match the CDR totals to take account of differences between logbook and landings data (multiple of 1.4 for the non-spawning fishery, based on $40 \%$ conversion from headed and gutted to whole, since 1986 and up to and including 1997 (reliable CDR data were available from 1998); 1.2 for the spawning fishery from 1986 up to and including 1996 (when factory vessels entered the spawning fishery) (D. Smith, pers. comm.). As stated by Thomson and He (2001), the factor is lower for the spawning fleet than the non-spawning fleet because some fish in the spawning fishery, landed headed and gutted, were recorded as being landed whole. These factors were chosen by the Blue Grenadier Assessment Group (BGAG) (Chesson and Staples (1995), as cited by Punt (1998)). The adjusted logbook catches were then scaled up to the SEF2 data (CDR). As historical CDR data were only available from 1992, the average scaling factor from 1992 to 1996 (1.07) was used to scale the data for years between 1986 and 1991. Note that in years 2008 to 2013 logbook data were greater than landings from the CDR. In these cases, the tonnage from the CDR was used as the total catch (AFMA, pers. comm. 2011). Table 6.2 lists the annual catches used in the assessment and the annual TAC (Figure 6.2). The annual logbook catches by sub-fishery and the adjustments made to determine the catches used in the assessment are shown in Table 6.1. No state catches are included and are assumed negligible or included in the historical values.

### 6.4.1.2 Discards

Discard rates were estimated from onboard data which gives the weight of the retained and discarded component of those shots that were monitored (Thomson and Klaer, 2011, Burch et al 2018). The discard rates are then scaled up to discard mass. The discard values from 1995 to 2002 are based on estimates calculated from ISMP data by MAFRI and reported in He et al. (1999) and Tuck, Smith and Talman (2004). The MAFRI estimates of discards were made accounting for differences in sampling and discard rates according to the ISMP zones. As agreed by Slope RAG (2011), since 2003 discard rates are estimated using the methods described in Thomson and Klaer (2011). Tier 1 stock assessments implemented in Stock Synthesis estimate discards within the assessment by fitting to discard proportions or mass calculated by fleet. Discard proportions are estimated for a population (stock) by fleet, year, zone and season (usually a quarter) and then scaled to landed (CDR) catch to obtain estimates by population, fleet and year (Klaer 2018). The discard proportion is estimated as the sum of the discarded catch divided by the sum of discarded catch and the landed catch (Klaer 2018; Method 1). The previous assessment used Method 2, where the discard proportion was estimated as the average of the proportion discarded in each shot (Klaer 2018). However, Method 2 does not scale the mean discard proportion by shot weight and it is therefore sensitive to the discarding practices from shots with small catches and, as such, may not be representative of the overall fishery. At its August 2020 Data Meeting SESSFRAG endorsed the use of Method 1 to estimate discard proportions for Tier 1 assessments from 2020 onwards. The discard rates calculated for and input to Tier 1 stock assessments are used to fit retention selectivity curves, so individual year values are not greatly influential on model estimated discard rates. Information in support of the historical values was not able to be obtained and further exploration of the methods and data used to estimate these values should be encouraged. The discard data are provided in Table 6.2. The discard data were assumed to have standard error (on the log-scale) of 0.3 . As with previous assessments, only discards from the non-spawning fishery are considered.


Figure 6.2. A comparison of total annual catches from the 2018 base case assessment and the updated catch used in the 2021 assessment for the spawning (S) and non-spawning (NS) fisheries.


Figure 6.3. A comparison of total annual estimated discard mass from the 2018 base case assessment and the updated catch used in the 2021 assessment for the non-spawning fishery.

Table 6.1. Logbook and CDR landings for the spawning and non-spawning sub-fisheries by calendar year and adjustments made to account for logbooks being less than landings and incorrect reporting process code. Shaded CDR are historical landings values. ${ }^{1}$ average of CDR/logbook ratio from 1992 to 1996.

| Year | Logbook |  | CDR | H\&G Multiplier |  | Adjusted Logbook |  | Total | CDR / <br> logbook | Catch for assessment |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Spawning | Non- |  | Spawnin | Non- | Spawning | Non- |  |  | Spawning | Non- |
| 1979 | 245 | 245 |  | 1 | 1 | 245 | 245 | 490 | 1.00 | 245 | 245 |
| 1980 | 410 | 410 |  | 1 | 1 | 410 | 410 | 820 | 1.00 | 410 | 410 |
| 1981 | 225 | 225 |  | 1 | 1 | 225 | 225 | 450 | 1.00 | 225 | 225 |
| 1982 | 390 | 390 |  | 1 | 1 | 390 | 390 | 780 | 1.00 | 390 | 390 |
| 1983 | 450 | 450 |  | 1 | 1 | 450 | 450 | 900 | 1.00 | 450 | 450 |
| 1984 | 675 | 675 |  | 1 | 1 | 675 | 675 | 1350 | 1.00 | 675 | 675 |
| 1985 | 600 | 600 |  | 1 | 1 | 600 | 600 | 1200 | 1.00 | 600 | 600 |
| 1986 | 246 | 1204 |  | 1.2 | 1.4 | 295 | 1685 | 1981 | 1.07 | 317 | 1806 |
| 1987 | 782 | 1455 |  | 1.2 | 1.4 | 939 | 2036 | 2975 | 1.07 | 1006 | 2183 |
| 1988 | 319 | 1485 |  | 1.2 | 1.4 | 383 | 2079 | 2461 | 1.07 | 410 | 2228 |
| 1989 | 36 | 1829 |  | 1.2 | 1.4 | 43 | 2560 | 2604 | 1.07 | 46 | 2745 |
| 1990 | 570 | 1671 |  | 1.2 | 1.4 | 684 | 2340 | 3023 | 1.07 | 733 | 2508 |
| 1991 | 637 | 2508 |  | 1.2 | 1.4 | 764 | 3511 | 4275 | 1.071 | 819 | 3764 |
| 1992 | 509 | 1565 | 3259 | 1.2 | 1.4 | 610 | 2191 | 2802 | 1.16 | 710 | 2549 |
| 1993 | 812 | 1659 | 3362 | 1.2 | 1.4 | 975 | 2323 | 3298 | 1.02 | 994 | 2368 |
| 1994 | 974 | 1338 | 3151 | 1.2 | 1.4 | 1169 | 1873 | 3042 | 1.04 | 1211 | 1940 |
| 1995 | 911 | 1017 | 2775 | 1.2 | 1.4 | 1093 | 1424 | 2517 | 1.10 | 1205 | 1570 |
| 1996 | 1200 | 1061 | 3040 | 1.2 | 1.4 | 1439 | 1485 | 2925 | 1.04 | 1496 | 1544 |
| 1997 | 2623 | 997 | 4516 | 1 | 1.4 | 2623 | 1396 | 4019 | 1.12 | 2947 | 1569 |
| 1998 | 2739 | 1459 | 5733 | 1 | 1 | 2739 | 1459 | 4198 | 1.37 | 3740 | 1993 |
| 1999 | 5460 | 2068 | 9324 | 1 | 1 | 5460 | 2068 | 7528 | 1.24 | 6762 | 2562 |
| 2000 | 5735 | 1761 | 8655 | 1 | 1 | 5735 | 1761 | 7496 | 1.15 | 6622 | 2033 |
| 2001 | 7309 | 1034 | 9128 | 1 | 1 | 7309 | 1034 | 8343 | 1.09 | 7997 | 1131 |
| 2002 | 6825 | 1151 | 9165 | 1 | 1 | 6825 | 1151 | 7976 | 1.15 | 7843 | 1322 |
| 2003 | 7239 | 687 | 8480 | 1 | 1 | 7239 | 687 | 7926 | 1.07 | 7746 | 735 |
| 2004 | 4647 | 1225 | 6401 | 1 | 1 | 4647 | 1225 | 5872 | 1.09 | 5066 | 1336 |
| 2005 | 2880 | 1204 | 4293 | 1 | 1 | 2880 | 1204 | 4085 | 1.05 | 3027 | 1266 |
| 2006 | 2058 | 1339 | 3625 | 1 | 1 | 2058 | 1339 | 3397 | 1.07 | 2196 | 1429 |
| 2007 | 1815 | 1232 | 3184 | 1 | 1 | 1815 | 1232 | 3048 | 1.04 | 1896 | 1287 |
| 2008 | 2838 | 1307 | 3938 | 1 | 1 | 2838 | 1307 | 4145 | 0.95 | 2696 | 1242 |


| 2098 |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 2009 | 2723 | 1151 | 3269 | 1 | 1 | 2723 | 1151 | 3874 | 0.84 | 2298 |
| 2010 | 3384 | 1162 | 4195 | 1 | 1 | 3384 | 1162 | 4545 | 0.92 | 3123 |
| 2011 | 3554 | 917 | 4207 | 1 | 1 | 3554 | 917 | 4471 | 0.94 | 3345 |
| 2012 | 3838 | 624 | 4063 | 1 | 1 | 3838 | 624 | 4461 | 0.91 | 3495 |
| 2013 | 3443 | 764 | 3828 | 1 | 1 | 3443 | 764 | 4207 | 0.91 | 3133 |
| 2014 | 279 | 935 | 1258 | 1 | 1 | 279 | 935 | 1215 | 1.04 | 289 |
| 2015 | 401 | 1061 | 1578 | 1 | 1 | 401 | 1061 | 1462 | 1.08 | 433 |
| 2016 | 217 | 978 | 1311 | 1 | 1 | 217 | 978 | 1195 | 1.10 | 238 |
| 2017 | 362 | 1261 | 1698 | 1 | 1 | 362 | 1261 | 1623 | 1.05 | 379 |
| 2018 | 508 | 1067 | 1665 | 1 | 1 | 508 | 1067 | 1575 | 1.06 | 537 |
| 2019 | 5799 | 1424 | 6914 | 1 | 1 | 5799 | 1424 | 7224 | 0.96 | 5551 |
| 2020 | 9146 | 1482 | 12151 | 1 | 1 | 9146 | 1482 | 10628 | 1.14 | 10457 |

Table 6.2. Landed and discarded catches for the spawning and non-spawning sub-fisheries by calendar year. These estimates have been scaled up to the landings data. Standardised CPUE (Sporcic, 2021) for the nonspawning sub-fisheries by calendar year are shown, along with the TAC. ${ }^{1}$ a voluntary industry reduction to $4,200 t$ was implemented in 2005. ${ }^{2}$ This was a 16 month TAC. ${ }^{3}$ From 2008/09, the TACs cover the fishing year 1 May to 30 April. In the table below, 2008 refers to 2008/09. * This is an estimate of retained catch equal to the 2020 catch.

| Year | Spawning (t) | Non-spawning (t) | Discards (t) | TAC | CPUE |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 1979 | 245 | 245 |  |  |  |
| 1980 | 410 | 410 |  |  |  |
| 1981 | 225 | 225 |  |  |  |
| 1982 | 390 | 390 |  |  |  |
| 1983 | 450 | 450 |  |  |  |
| 1984 | 675 | 675 |  |  |  |
| 1985 | 600 | 600 |  |  |  |
| 1986 | 317 | 1806 |  |  | 1.5312 |
| 1987 | 1006 | 2183 |  |  | 1.9494 |
| 1988 | 410 | 2228 |  |  | 2.1329 |
| 1989 | 46 | 2745 |  |  | 2.1313 |
| 1990 | 733 | 2508 |  |  | 2.1103 |
| 1991 | 819 | 3764 |  |  | 1.5098 |
| 1992 | 710 | 2549 |  |  | 1.2214 |
| 1993 | 994 | 2368 |  |  | 0.9287 |
| 1994 | 1211 | 1940 |  | 10000 | 0.8412 |
| 1995 | 1205 | 1570 | 80 | 10000 | 0.5802 |
| 1996 | 1496 | 1544 | 975 | 10000 | 0.5262 |
| 1997 | 2947 | 1569 | 3716 | 10000 | 0.5464 |
| 1998 | 3740 | 1993 | 1329 | 10000 | 0.8818 |
| 1999 | 6762 | 2562 | 123 | 10000 | 0.9257 |
| 2000 | 6622 | 2033 | 69 | 10000 | 0.6643 |
| 2001 | 7997 | 1131 | 10 | 10000 | 0.3828 |
| 2002 | 7843 | 1322 | 2 | 10000 | 0.3794 |
| 2003 | 7746 | 735 | 16 | 9000 | 0.3171 |
| 2004 | 5066 | 1336 | 35 | 7000 | 0.5326 |
| 2005 | 3027 | 1266 | 275 | $5000^{1}$ | 0.6428 |
| 2006 | 2196 | 1429 | 91 | 3730 | 0.8564 |
| 2007 | 1896 | 1287 | 40 | $4113^{2}$ | 0.7622 |
| 2008 | 2696 | 1242 | 36 | $4368^{3}$ | 0.8386 |
| 2009 | 2298 | 971 | 76 | 4700 | 0.7778 |
| 2010 | 3123 | 1072 | 56 | 4700 | 0.7805 |
| 2011 | 3345 | 863 | 123 | 4700 | 0.637 |
| 2012 | 3495 | 568 | 281 | 5208 | 0.508 |
| 2013 | 3133 | 695 | 311 | 5208 | 0.9059 |
| 2014 | 289 | 969 | 455 | 6800 | 1.092 |
| 2015 | 433 | 1146 | 601 | 8796 | 1.1867 |
| 2016 | 238 | 1073 | 619 | 8810 | 1 |
| 2017 | 379 | 1319 | 576 | 8765 | 1.1183 |
| 2018 | 537 | 1128 | 317 | 8810 | 0.899 |
| 2019 | 5551 | 1363 | 659 | 12183 | 1.1917 |
| 2020 | 10457 | 1694 | 598 | 12183 | 1.7107 |
| 2021 | 10457* | 1694* |  |  |  |

### 6.4.2 Catch rates

Sporcic (2021) provides the updated standardised catch rate series for the non-spawning fishery of Blue Grenadier (Table 6.2; Figure 6.4). The catch rate generally follows the fluctuations of stock size driven by large, but sporadic, recruitments. The standard deviation of log-CPUE is assumed to be 0.252 (value equal to the standard error from a loess fit), but an extra variance component is estimated for the CPUE index during the tuning process.


Figure 6.4. A comparison of the annual standardised catch rates series for Blue Grenadier between the 2018 and 2021 assessments.

### 6.4.3 Length-composition and age data

Length and age data are included in the assessment as length-composition data and conditional age-atlength data by fleet and sex (the latter if available). Onboard and port length-compositions, when available, are used separately. Separating port and onboard lengths first occurred in the 2018 assessment. Prior to 2018, only port samples had been used to create the length-compositions. Plots of the observed length and age data are shown in later figures, with the corresponding model predicted values.

There had to be at least 100 measured fish for a retained and/or discard onboard and port lengthcomposition data to be included in the assessment. For onboard samples, numbers of shots were used as the sampling unit (i.e. the stage-1 weights; Francis, 2011), with a cap of 200. For port samples, numbers of trips were used as the sampling unit, with a cap of 100 . The number of fish measured is not used as the sample size because the appropriate sample size for length-composition data is probably more closely related to the number of shots (onboard) or trips (port) sampled, rather than the number of fish measured (Table 6.3; Table 6.4).

Table 6.3. The years for which length data were available for the sub-fleets (spawning onboard $=1$; spawning port $=3$; non-spawning onboard $=2$; non-spawning port $=4)$, $\operatorname{sex}(0=$ no gender specified; female $=1$; male $=2$ ), partition (part: discard $=1$; retained $=2$ ). N is the number of shots (onboard) or trips (port). Red length data were excluded due to low sample sizes. ${ }^{1}$ the average number of fish from years 1984 and 1988. ${ }^{2}$ these years of discard lengths were removed due to spurious numbers of large fish.

| Year | Nfish | Fleet | Sex | Part | N |
| ---: | ---: | ---: | ---: | ---: | ---: |
| 1984 | 1046 | 1 | 0 | 2 | 12 |
| 1985 | $1090^{1}$ | 1 | 0 | 2 | 12 |
| 1988 | 1133 | 1 | 0 | 2 | 12 |
| 1998 | 812 | 1 | 0 | 2 | 10 |
| 1998 | 1037 | 1 | 1 | 2 | 8 |
| 1998 | 469 | 1 | 2 | 2 | 8 |
| 1999 | 4147 | 1 | 1 | 2 | 79 |
| 1999 | 5929 | 1 | 2 | 2 | 79 |
| 2000 | 2672 | 1 | 1 | 2 | 48 |
| 2000 | 2956 | 1 | 2 | 2 | 46 |
| 2001 | 3620 | 1 | 1 | 2 | 67 |
| 2001 | 4256 | 1 | 2 | 2 | 67 |
| 2002 | 262 | 1 | 0 | 2 | 2 |
| 2002 | 444 | 1 | 1 | 2 | 3 |
| 2002 | 450 | 1 | 2 | 2 | 3 |
| 2003 | 2700 | 1 | 1 | 2 | 59 |
| 2003 | 2853 | 1 | 2 | 2 | 59 |
| 2004 | 1307 | 1 | 1 | 2 | 28 |
| 2004 | 1370 | 1 | 2 | 2 | 28 |
| 2005 | 198 | 1 | 1 | 2 | 20 |
| 2005 | 141 | 1 | 2 | 2 | 20 |
| 2006 | 3184 | 1 | 1 | 2 | 56 |
| 2006 | 3081 | 1 | 2 | 2 | 55 |
| 2007 | 2957 | 1 | 1 | 2 | 54 |
| 2007 | 1897 | 1 | 2 | 2 | 55 |
| 2008 | 3073 | 1 | 1 | 2 | 53 |
| 2008 | 2177 | 1 | 2 | 2 | 54 |
| 2009 | 3868 | 1 | 1 | 2 | 73 |
| 2009 | 3374 | 1 | 2 | 2 | 70 |
| 2010 | 2488 | 1 | 1 | 2 | 98 |
| 2010 | 1453 | 1 | 2 | 2 | 94 |
| 2011 | 4207 | 1 | 1 | 2 | 79 |
| 2011 | 3266 | 1 | 2 | 2 | 77 |
| 2012 | 3939 | 1 | 1 | 2 | 77 |
| 2012 | 3060 | 1 | 2 | 2 | 82 |
| 2013 | 1 | 1 | 0 | 2 | 1 |
| 2013 | 4443 | 1 | 1 | 2 | 76 |
| 2013 | 3892 | 1 | 2 | 2 | 76 |
| 2014 | 592 | 1 | 0 | 2 | 7 |
| 2014 | 229 | 1 | 1 | 2 | 9 |
| 2014 | 179 | 1 | 2 | 2 | 9 |
| 2015 | 715 | 1 | 0 | 2 | 11 |
|  |  |  |  |  |  |


| 2015 | 723 | 1 | 1 | 2 | 18 |
| ---: | ---: | ---: | ---: | ---: | ---: |
| 2015 | 862 | 1 | 2 | 2 | 18 |
| 2017 | 777 | 1 | 0 | 2 | 12 |
| 2017 | 131 | 1 | 1 | 2 | 11 |
| 2017 | 193 | 1 | 2 | 2 | 11 |
| 2018 | 10 | 1 | 0 | 2 | 1 |
| 2019 | 57 | 1 | 0 | 2 | 19 |
| 2019 | 3389 | 1 | 1 | 2 | 72 |
| 2019 | 4324 | 1 | 2 | 2 | 72 |
| 2020 | 8 | 1 | 0 | 2 | 6 |
| 2020 | 6776 | 1 | 1 | 2 | 204 |
| 2020 | 8774 | 1 | 2 | 2 | 201 |


| Year | Nfish | Fleet | Sex | Part | N |
| ---: | ---: | ---: | ---: | ---: | ---: |
| 1984 | 1935 | 2 | 0 | 2 | 75 |
| 1985 | 1829 | 2 | 0 | 2 | 99 |
| 1987 | 4063 | 2 | 0 | 2 | 100 |
| 1988 | 6660 | 2 | 0 | 2 | 164 |
| 1989 | 2424 | 2 | 0 | 2 | 160 |
| 1996 | 829 | 2 | 0 | 2 | 8 |
| 1997 | 3367 | 2 | 0 | 2 | 32 |
| 1998 | 8290 | 2 | 0 | 2 | 73 |
| 1999 | 8768 | 2 | 0 | 2 | 79 |
| 2000 | 9362 | 2 | 0 | 2 | 73 |
| 2001 | 6309 | 2 | 0 | 2 | 57 |
| 2002 | 5329 | 2 | 0 | 2 | 47 |
| 2003 | 2754 | 2 | 0 | 2 | 50 |
| 2004 | 7586 | 2 | 0 | 2 | 104 |
| 2005 | 5754 | 2 | 0 | 2 | 76 |
| 2006 | 6549 | 2 | 0 | 2 | 68 |
| 2007 | 1109 | 2 | 0 | 2 | 44 |
| 2008 | 2624 | 2 | 0 | 2 | 91 |
| 2009 | 2100 | 2 | 0 | 2 | 79 |
| 2010 | 2562 | 2 | 0 | 2 | 71 |
| 2011 | 1755 | 2 | 0 | 2 | 70 |
| 2012 | 3087 | 2 | 0 | 2 | 97 |
| 2013 | 1841 | 2 | 0 | 2 | 48 |
| 2014 | 2631 | 2 | 0 | 2 | 67 |
| 2015 | 1555 | 2 | 0 | 2 | 45 |
| 2016 | 3960 | 2 | 0 | 2 | 68 |
| 2017 | 1236 | 2 | 0 | 2 | 18 |
| 2018 | 1585 | 2 | 0 | 2 | 38 |
| 2019 | 2579 | 2 | 0 | 2 | 53 |
| 2020 | 1261 | 2 | 0 | 2 | 33 |
|  |  |  |  |  |  |


| Year | Nfish | Fleet | Sex | Part | N |
| :---: | :---: | :---: | :---: | :---: | :---: |
| $1992{ }^{2}$ | 159 | 2 | 0 | 1 | 3 |
| $1993{ }^{2}$ | 1532 | 2 | 0 | 1 | 12 |
| $1994{ }^{2}$ | 2366 | 2 | 0 | 1 | 27 |
| $1995{ }^{2}$ | 6651 | 2 | 0 | 1 | 61 |
| $1996{ }^{2}$ | 5999 | 2 | 0 | 1 | 50 |
| 1997 | 6967 | 2 | 0 | 1 | 62 |
| 1998 | 2212 | 2 | 0 | 1 | 20 |
| 1999 | 940 | 2 | 0 | 1 | 7 |
| 2000 | 132 | 2 | 0 | 1 | 3 |
| 2003 | 11 | 2 | 0 | 1 | 6 |
| 2004 | 1078 | 2 | 0 | 1 | 22 |
| 2005 | 5299 | 2 | 0 | 1 | 48 |
| 2006 | 1225 | 2 | 0 | 1 | 8 |
| 2007 | 16 | 2 | 0 | 1 | 2 |
| 2008 | 219 | 2 | 0 | 1 | 18 |
| 2009 | 97 | 2 | 0 | 1 | 6 |
| 2010 | 16 | 2 | 0 | 1 | 2 |
| 2011 | 792 | 2 | 0 | 1 | 30 |
| 2012 | 1327 | 2 | 0 | 1 | 49 |
| 2013 | 1455 | 2 | 0 | 1 | 41 |
| 2014 | 873 | 2 | 0 | 1 | 17 |
| 2015 | 500 | 2 | 0 | 1 | 18 |
| 2016 | 1360 | 2 | 0 | 1 | 28 |
| 2017 | 531 | 2 | 0 | 1 | 9 |
| 2018 | 682 | 2 | 0 | 1 | 13 |
| 2019 | 151 | 2 | 0 | 1 | 8 |
| 2020 | 32 | 2 | 0 | 1 | 5 |
| 1992 | 774 | 3 | 0 | 2 | 6 |
| 1994 | 1038 | 3 | 0 | 2 | 9 |
| 1995 | 465 | 3 | 0 | 2 | 4 |
| 1996 | 927 | 3 | 0 | 2 | 7 |
| 1997 | 851 | 3 | 0 | 2 | 7 |
| 1998 | 1648 | 3 | 0 | 2 | 9 |
| 1999 | 1079 | 3 | 0 | 2 | 9 |
| 2000 | 360 | 3 | 0 | 2 | 3 |
| 2014 | 82 | 3 | 0 | 2 | 1 |
| 2016 | 74 | 3 | 0 | 2 | 1 |
| 2020 | 100 | 3 | 0 | 2 | 1 |


| Year | Nfish | Fleet | Sex | Part | N |
| ---: | ---: | ---: | ---: | ---: | ---: |
| 1979 | 164 | 4 | 0 | 2 | 2 |
| 1980 | 40 | 4 | 0 | 2 | 1 |
| 1981 | 1425 | 4 | 0 | 2 | 36 |
| 1982 | 478 | 4 | 0 | 2 | 12 |
| 1991 | 927 | 4 | 0 | 2 | 10 |
| 1992 | 3832 | 4 | 0 | 2 | 31 |
| 1993 | 1810 | 4 | 0 | 2 | 12 |
| 1994 | 8624 | 4 | 0 | 2 | 79 |
| 1995 | 7055 | 4 | 0 | 2 | 62 |
| 1996 | 5505 | 4 | 0 | 2 | 51 |
| 1997 | 11844 | 4 | 0 | 2 | 85 |
| 1998 | 16234 | 4 | 0 | 2 | 100 |
| 1999 | 13898 | 4 | 0 | 2 | 119 |
| 2000 | 13728 | 4 | 0 | 2 | 95 |
| 2001 | 12000 | 4 | 0 | 2 | 88 |
| 2002 | 9416 | 4 | 0 | 2 | 77 |
| 2003 | 5037 | 4 | 0 | 2 | 38 |
| 2004 | 4440 | 4 | 0 | 2 | 43 |
| 2005 | 6310 | 4 | 0 | 2 | 48 |
| 2006 | 3019 | 4 | 0 | 2 | 31 |
| 2007 | 979 | 4 | 0 | 2 | 9 |
| 2008 | 1955 | 4 | 0 | 2 | 16 |
| 2009 | 1080 | 4 | 0 | 2 | 19 |
| 2010 | 833 | 4 | 0 | 2 | 26 |
| 2011 | 1925 | 4 | 0 | 2 | 54 |
| 2012 | 1331 | 4 | 0 | 2 | 33 |
| 2013 | 1744 | 4 | 0 | 2 | 43 |
| 2014 | 1611 | 4 | 0 | 2 | 30 |
| 2015 | 2048 | 4 | 0 | 2 | 25 |
| 2016 | 1887 | 4 | 0 | 2 | 29 |
| 2017 | 2061 | 4 | 0 | 2 | 35 |
| 2018 | 1943 | 4 | 0 | 2 | 27 |
| 2019 | 1222 | 4 | 0 | 2 | 22 |
| 2020 | 1864 | 4 | 0 | 2 | 32 |
|  |  |  |  |  |  |

Table 6.4. Number of age-length otolith samples by fleet included in the base case assessment.

| Year | Spawn | Non-spawn |
| ---: | ---: | ---: |
| 1984 | 512 | 735 |
| 1985 | 432 | 603 |
| 1986 | 174 | 71 |
| 1987 |  | 1027 |
| 1988 |  | 1092 |
| 1989 |  | 1031 |
| 1990 |  |  |
| 1991 | 93 | 100 |
| 1992 | 481 | 706 |
| 1993 | 1122 | 772 |
| 1994 | 1130 | 623 |
| 1995 | 1154 | 637 |
| 1996 | 1296 | 932 |
| 1997 | 932 | 1697 |
| 1998 | 1334 | 948 |
| 1999 | 992 | 802 |
| 2000 | 1247 | 1224 |
| 2001 | 1062 | 891 |
| 2002 | 1077 | 751 |
| 2003 | 1035 | 514 |
| 2004 | 1187 | 435 |
| 2005 | 1016 | 1185 |
| 2006 | 1313 | 816 |
| 2007 | 1205 | 396 |
| 2008 | 1437 | 753 |
| 2009 | 1545 | 907 |
| 2010 | 1530 | 451 |
| 2011 | 1515 | 763 |
| 2012 | 1391 | 715 |
| 2013 | 1655 | 621 |
| 2014 | 884 | 887 |
| 2015 | 696 | 723 |
| 2016 | 221 | 773 |
| 2017 | 537 | 928 |
| 2018 | 221 | 733 |
| 2019 | 1406 | 1119 |
| 2020 | 1579 | 344 |
|  |  |  |

### 6.4.4 Acoustic survey estimates

Estimates of spawning biomass for 2003-2010 are provided in Ryan and Kloser (2012). There are no acoustic estimates since 2010. Table 6.5 shows the estimates of spawning biomass with their corresponding CV's used in the assessment. Sampling CVs less than 0.3 were increased to 0.3 to account for process error. Low sampling CVs (of 0.19 for example) were considered too low for an acoustic survey and a minimum of 0.3 should be used to reflect the total uncertainty (D. Smith, pers comm., Tuck et al., 2004; Slope RAG 2011). Of 22 acoustic CVs used for Hoki in New Zealand, none are lower than 0.3 (Francis, 2009). It is assumed that the spawning ground experiences a turnover rate of two (i.e. for the model applied here, the spawning biomass estimates are doubled) (Russell and Smith, 2006; Punt et al., 2015). The acoustic survey selectivity is matched to the maturity ogive, as it is assumed the acoustic survey observes mature fish on the spawning ground.

Table 6.5. The estimated biomass (tonnes) of Blue Grenadier on the spawning grounds in years 2003 to 2010 (Ryan and Kloser, 2012).

|  | 2003 | 2004 | 2005 | 2006 | 2007 | 2008 | 2009 | 2010 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Biomass (t) <br> CV for <br> assessment <br> model | 24,690 | 16,295 | 18,852 | 42,882 | 56,330 | 24,450 | 24,787 | 20,622 |
| Sampling CV <br> Sam | 0.30 | 0.46 | 0.30 | 0.30 | 0.52 | 0.30 | 1 | 0.33 |

### 6.4.5 Egg survey estimates

Egg survey estimates of female spawning biomass are available for 1994 and 1995 (Bulman et al., 1999). The egg-estimates (CV) for 1994 and 1995 respectively are: 57,772 (0.18) and 41,409 (0.29) tonnes. For the analysis considered here, the base-case egg estimates were used.

### 6.4.6 Biological parameters and stock structure assumptions

The assessment assumes that the proportion of females that spawn in each year is 0.84 and a length at $50 \%$ maturity of 63.7 cm for females (Russel and Smith, 2006). The female maturity ogive is shown in Figure 6.4.

The length weight-relationship for males and females was estimated from spawning fishery data over years 1999 to 2008 (Figure 6.5). Natural mortality for females and males is estimated when fitting the model.

Francis (2009) reviews the values of steepness used in New Zealand Hoki assessments, where a value of $h=0.9$ had been used since 1994. This value of steepness was derived from work of Punt et al. (1994) using 45 stocks of Gadiform species ( 0.9 is the median). Following an analysis of the profile likelihood, the effect of steepness on the 2007 assessment and additional information of Myers et al. (1999; 2002) beyond that used by Punt et al. (1994), Francis (2009) concludes that steepness should be reduced to $h=0.75$. This value of steepness has been assumed in all Blue Grenadier assessments since 2011 and in this assessment.


Figure 6.5. The maturity ogive by length for female Blue Grenadier (parameters from Russell and Smith (2006)) and the length-weight relationship for males and females.

### 6.4.7 Age-reading error

Updated standard deviations for aging error by reader (A and B) have been estimated, producing the age-reading error matrix of Table 6.6 (A. Punt and P. Burch, pers. comm.). Reader A applied to years 1991-93 and 2007-20, and reader B to years 1984-90 and 1994-2006.

Table 6.6. The standard deviation of age reading error for readers A and B.

| St Dev |  |  |
| :---: | :---: | :---: |
| Age | A | B |
| 0 | 0.198 | 0.281 |
| 1 | 0.198 | 0.281 |
| 2 | 0.258 | 0.299 |
| 3 | 0.305 | 0.318 |
| 4 | 0.341 | 0.338 |
| 5 | 0.369 | 0.359 |
| 6 | 0.391 | 0.383 |
| 7 | 0.407 | 0.408 |
| 8 | 0.420 | 0.435 |
| 9 | 0.430 | 0.464 |
| 10 | 0.438 | 0.495 |
| 11 | 0.444 | 0.529 |
| 12 | 0.448 | 0.565 |
| 13 | 0.452 | 0.604 |
| 14 | 0.455 | 0.646 |
| 15 | 0.457 | 0.691 |
| 16 | 0.459 | 0.740 |
| 17 | 0.460 | 0.792 |
| 18 | 0.461 | 0.848 |
| 19 | 0.462 | 0.908 |
| 20 | 0.462 | 0.974 |

### 6.5 Analytical Approach

### 6.5.1 Model structure and parameters

The 2021 base case assessment of Blue Grenadier uses an age- and size-structured model implemented in the generalized stock assessment software package, Stock Synthesis (SS) (Version 3.30.17.00, Methot et al. (2021)). The methods utilised in SS are based on the integrated analysis paradigm. SS can allow for multiple seasons, areas and fleets, but most applications are based on a single season and area. Recruitment is governed by a stochastic Beverton-Holt stock-recruitment relationship, parameterized in terms of the steepness of the stock-recruitment function (h), the expected average recruitment in an unfished population ( $R_{0}$ ), and the degree of variability about the stock-recruitment relationship $\left(\sigma_{r}\right)$. SS allows the user to choose among a large number of age- and length-specific selectivity patterns. The values for the parameters of SS are estimated by fitting to data on catches, catch-rates, discard mass, discard and retained catch length-frequencies, and conditional age-at-length data. The population dynamics model and the statistical approach used in fitting the model to the various data types are given in the SS technical documentation.

Model data have been updated by the inclusion of data up to the 2020 calendar year (lengthcomposition and conditional age-at-length data; age reading-error matrices, standardized catch rate series; landings and discard catch weight) and information from acoustic surveys of spawning biomass (series from 2003-2010, pertaining to total spawning biomass), with an assumption of two-times turnover on the spawning ground (Russell and Smith, 2006; Punt et al. 2015). The base-case egg survey estimates of female (only) spawning biomass for 1994 and 1995 are included. The model fits directly to length-composition data (by sex where possible) and conditional age-at-length data by fleet. Retained length-composition data from port and onboard samples are separated.

The base-case model includes the following key features:
a) Blue grenadier consists of a single stock within the area of the fishery.
b) The model accounts for males and females separately (growth, natural mortality, age at first breeding).
c) The population was at its unfished biomass with the corresponding equilibrium (unfished) agestructure at the start of 1960 .
d) The rate of natural mortality, $M$, is assumed to be constant with age, and also time-invariant. The value for female and male $M$ is estimated within the assessment.
e) Recruitment to the stock is assumed to follow a Beverton-Holt type stock-recruitment relationship, parameterised by the average recruitment at unexploited spawning biomass, $R_{0}$, and the steepness parameter, $h$. Steepness for the base-case analysis is set to 0.75 . Deviations from the average recruitment at a given spawning biomass (recruitment residuals) are estimated for 1974 to 2017. Deviations are not estimated before 1974 or after 2017 because there are insufficient data to permit reliable estimation of recruitment residuals outside of this time period.
f) The population plus-group is modelled at age 20 years. The maximum age for age observations is 20 years.
g) Growth is assumed to follow a von Bertalanffy type length-at-age relationship, with the parameters of the growth function being estimated separately for females and males inside the assessment model. Growth is also assumed to vary through time and to be cohort (year class) specific. Evidence for time-varying and cohort specific growth in Blue Grenadier has been accumulating over several decades (see Whitten et al., 2013). The 2021 base-case model treats
conditional age-at-length information as data (i.e. to incorporate error), and predicts the expected length-at-age for each year. This is achieved by estimating the parameters of a von Bertalanffy growth function where the expected annual growth increment is based on the von Bertalanffy growth function but with a growth rate parameter that is determined by an expected value and a cohort-specific deviation. Cohort-specific deviations from average growth are estimated in the base case model for year classes 1978 to 2017.
h) Two fleets are included in the model - the spawning fishery that operates during winter (June August inclusive) off western Tasmania (zone 40), and the non-spawning sub-fishery that operates during other times of the year and in other areas throughout the year. GAB catches are not included.
i) Each selectivity pattern was assumed to be length-specific, logistic and time-invariant for the spawning fleet and dome-shaped for the non-spawning fleet. The parameters of the selectivity function for each fleet were estimated within the assessment.
j) The CVs of the CPUE indices were initially set at a value equal to the standard error from a loess fit ( 0.252 ; Sporcic, 2021), before being re-tuned to the model-estimated standard errors within SS. The acoustic estimates were tuned through the estimation of an extra variance component that is added to the model input standard errors. This is done within SS.
k) Discard tonnage was estimated through the assignment of a retention function for the nonspawning fleet. This was defined as a logistic function of length, and the inflection and slope of this function were estimated where discard information was available. In addition, the discard length data from prior to 1996 were removed as recommended by SERAG (September, 2018) due to the existence of unusually large fish in the length distribution which is likely to be misreporting.

1) Retained and discarded onboard length sample sizes were capped at 200 and a minimum of 100 fish measured was required for length-composition data to be included in the assessment. For port samples, numbers of trips were used as the sampling unit, with a cap of 100 . The number of fish measured is not used as the sample size because the appropriate sample size for lengthcomposition data is probably more closely related to the number of shots (onboard) or trips (port) sampled, rather than the number of fish measured.

The values assumed for fixed parameters of the preliminary base case model are shown in Table 6.7.

Table 6.7. Parameter values assumed for some of the non-estimated parameters of the base-case model

| Parameter | Description | Value |
| :---: | :---: | :---: |
| $M_{f}$ | Natural mortality for females | Estimated |
| $M_{m}$ | Natural mortality for males | Estimated |
| h | "steepness" of the Beverton-Holt stock-recruit curve | 0.75 |
| x | age observation plus group | 20 years |
| $\mu$ | fraction of mature population that spawn each year | 0.84 |
| $a_{f}$ | Female allometric length-weight equations | $0.01502 \mathrm{~g}^{-1} \mathrm{~cm}$ |
| $b_{f}$ | Female allometric length-weight equations | 2.728 |
| $a_{m}$ | Male allometric length-weight equations | $0.0168 \mathrm{~g}^{-1} \mathrm{~cm}$ |
| $b_{m}$ | Male allometric length-weight equations | 2.680 |
| $l_{m}$ | Female length at 50\% maturity | 63.7 cm |
| $l_{s}$ | Parameter defining the slope of the maturity ogive | -0.261 |

### 6.5.2 Tuning Method

Iterative rescaling (reweighting) of input and output CVs or input and effective sample sizes is a repeatable method for ensuring that the expected variation of the different data streams is comparable to what is input (Pacific Fishery Management Council, 2018). Most of the indices (CPUE, surveys and composition data) used in fisheries underestimate their true variance by only reporting measurement or estimation error and not including process error.

In iterative reweighting, the effective annual sample sizes are tuned/adjusted so that the input sample size is equal to the effective sample size calculated by the model. In SS-V3.30 it is possible to estimate an additional standard deviation parameter to add to the input CVs for the abundance indices (CPUE).

1. Set the standard error for the log of relative abundance indices (CPUE) to the standard deviation of a loess curve fitted to the original data - which will provide a more realistic estimate to that obtained from the original statistical analysis. SS-V3.30 then allows an estimate to be made for an additional adjustment to the relative abundance variances appropriately.

An automated iterative tuning procedure was used for the remaining adjustments. For the recruitment bias adjustment ramps:
2. Adjust the maximum bias adjustment and the start and finish bias adjustment ramps as predicted by SS-V3.30 at each step.

For the age and length composition data:
3. Multiply the stage-1 (initial) sample sizes for the conditional age-at-length data by the sample size multipliers using the approach of Punt (2017).
4. Similarly multiply the initial samples sizes by the sample size multipliers for the length composition data using the 'Francis method' (Francis, 2011).
5. Repeat steps $2-4$, until all are converged and stable (with proposed changes $<1-2 \%$ ).

This procedure constitutes current best practice for tuning assessments.

### 6.5.3 Calculating the RBC

The SESSF Harvest Strategy Framework (HSF) was developed during 2005 (Smith et al., 2008) and has been used as a basis for providing advice on TACs in the SESSF quota management system for fishing years 2006-2020. The HSF uses harvest control rules to determine a recommended biological catch (RBC) for each stock in the SESSF quota management system. Each stock is assigned to a Tier level depending on the basis used for assessing stock status or exploitation level for that stock. Blue Grenadier is assessed as a Tier 1 stock as it has an agreed quantitative stock assessment.

The Tier 1 harvest control rule specifies a target and a limit biomass reference point, as well as a target fishing mortality rate. Since 2005 various values have been used for the target and the breakpoint in the rule. The 20:40:40 ( $B_{\mathrm{lim}}: B_{\mathrm{MSY}}: F_{\mathrm{targ}}$ ) form of the rule is used up to where fishing mortality reaches $F_{48}$. Once this point is reached, the fishing mortality is set at $F_{48}$. Day (2008) has determined that for most SESSF stocks where the proxy values of $B_{40}$ and $B_{48}$ are used for $B_{\text {MSY }}$ and $B_{\text {MEY }}$ this form of the rule is equivalent to a 20:35:48 strategy.

This document reports RBCs calculated under the 20:35:48 strategy.

### 6.5.4 Sensitivity tests

A number of tests were used to examine the sensitivity of the results of the model to some of the assumptions and data inputs:

1. $h=0.85,0.65(0.75$ in the base case $)$
2. $M_{\text {fem }}=0.21,0.25(0.23$ in the base case $)$
3. Double and halve the weighting on the length composition data.
4. Double and halve the weighting on the age-at-length data.
5. Double and halve the weighting on the index (survey) data.
6. $\sigma_{r}=0.6,0.8$ ( 0.7 in the base case)

The results of the sensitivity tests are summarized by the following quantities:

1. $S B_{0}$ the average equilibrium female spawning biomass.
2. $S B_{2022}$ the female spawning biomass at the start of 2022.
3. $S B_{2022} / S B_{0}$ the depletion level at the start of 2022 , i.e. the 2022 spawning biomass expressed as a fraction of the unexploited spawning biomass.
4. $2022 R B C$ - the 2022 RBC , calculated using the $20: 35: 48$ harvest rule (presented for the agreed base case only).
5. Long-term RBC - the long-term RBC calculated using the 20:35:48 harvest rule (presented for the agreed base case only).

### 6.6 Results

### 6.6.1 The base-case analysis

### 6.6.1.1 Transition from the 2018 base case to the 2021 base case

The development of a preliminary base case, and a bridging analysis from the 2018 assessment (Castillo-Jordán and Tuck, 2018b), was presented at the October 2021 SERAG 2 meeting (Tuck and Bessell-Browne, 2021), including updating the version of Stock Synthesis and sequentially updating data. This bridging analysis is not repeated in this report.

### 6.6.1.2 Paramater estimates

Figure 6.6 shows how the expected mean length-at-age values change over time for the base case model. The ridges reflect the impact of the estimated cohort dependent growth with some cohorts growing faster or slower than average. This figure also shows the expected mean length-at-age values for the end-year of the model. The impact of slower than average growth is visible by the decrease in expected size of say 10 year old fish in 2005, corresponding to the larger than average recruitment in 1994. Natural mortality for females was estimated to be $M_{f}=0.23$ and males was $M_{m}=0.24$.

The selectivity for the spawning and non-spawning fisheries and the retention function for the nonspawning fishery are shown in Figure 6.7. Selectivity is assumed to be time-invariant, sex-specific and logistic for the spawning fleet and dome-shaped for the non-spawning fleet.

The estimate of the parameter that defines the initial numbers (and biomass), $\ln (R 0)$, is 9.89 for the base case.


Figure 6.6. The estimate growth curve, with cohort dependent growth for Blue Grenadier.


Figure 6.7. Estimated selectivity for the spawning and non-spawning fleets, port and onboard samples and for males (m) and females (f) and the estimated retention function for the non-spawning fleet.

### 6.6.1.3 Fits to the data

Figure 6.8 shows the model fit to the non-spawning catch rate series. The model fits intersect most of the $95 \%$ confidence intervals for the data, indicating that adjustments to the CVs for the indices performed as expected. As has been seen in all previous assessment models for Blue Grenadier, the model is not able to fit the rise in catch rate following the large recruitment of the mid-1990s. More recent increases in catch rate are estimated well. The fit to the discard mass is able to replicate the increase in discarding through the late 1990s, mid-2000s and since 2012, however the magnitude is under-estimated (as has been the case with previous assessments). In the past, alternative models that time-blocked discarding, re-weighted discard CVs and included a discard fleet have all been unsuccessful in improving the fit to the discard and CPUE data. Further consideration should be given to the GLM model structure used in the standardisation of CPUE. Fits to the biomass estimates from the acoustic surveys and egg surveys were reasonable. The predicted biomass trajectory intersects all $95 \%$ confidence intervals.

The base-case model fits to the aggregated retained and discarded length-frequency distributions well (Figure 6.9). Note that a single selectivity is estimated for the combined port and onboard fleets. The saw-tooth port lengths which occurs when lengths measured in dorsal standard length (DSL), with values across all length bins, are converted to standard (STD) length, resulting in some length bins with lower estimates and higher estimates in neighbouring bins in the new length composition. Length composition fits by year and fleet are in the Appendix.


Figure 6.8. Fits to the non-spawning CPUE index, discard mass, egg survey and acoustic survey.

Length comps, aggregated across time by fleet


Figure 6.9. Length composition fits aggregated across years.

### 6.6.1.4 Assessment outcomes - base case

The estimated time series of recruitment under the base-case parameter set shows the typical episodic nature of Blue Grenadier recruitment, with strong year-classes in 1979, the mid-1980s, 1994, 2003, and from 2010 to 2017 (Figure 6.10). The trajectories of spawning biomass and spawning biomass relative to the un-exploited level are shown in Figure 6.10. This shows the increases and decreases in spawning biomass as the strong cohorts move into and out of the spawning population. Spawning biomass has varied considerably, with biomass below the target in 2013 and 2014, but nearly double virgin biomass in 1991, 2001 and 2021. Figure 6.11 shows various recruitment diagnostics and the annual recruitment deviations for the base case model. The figure showing recruitment deviations illustrates the historical episodic nature of recruitment, but also that the last eight estimates of recruitment are well above average. The Kobe plot in Figure 6.12 shows that the stock is well above
virgin biomass levels, but also that there is considerable uncertainty regarding both relative fishing mortality and stock status.

The estimated virgin female biomass is $37,445 \mathrm{t}$ (compared to $53,909 \mathrm{t}$ in 2018 and $36,815 \mathrm{t}$ in the 2013 assessments). Initial biomass is known to be sensitive in this model and often has varied betweem $35,000 \mathrm{t}$ and 60,000 t (Figure 6.13; Castillo- Jordán and Tuck, 2018a). A likelihood profile on initial biomass illustrates this uncertainty (Section 5.2).

For the base case model, the projected 2022 spawning stock biomass will be $155 \%$ of virgin female spawning biomass (projected assuming 2020 catches in 2021), compared to $122 \%$ for 2019 in the 2018 assessment, and $94 \%$ for 2014 in the 2013 assessment. The 2022 recommended biological catch (RBC) under the 20:35:48 harvest control rule is $23,777 \mathrm{t}$, with 245 t estimated discards ( $23,532 \mathrm{t}$ retained). The long-term RBC is $7,100 \mathrm{t}$, with 183 t discards (Table 6.8).

Table 6.8. The estimated RBC (tonnes), retained portion of the RBC, estimated discards and relative stock status for Blue Grenadier under the base case model. The retained catch up to 2020 is the actual tonnage (and 2021 catches are projected assuming 2020 catches in 2021), and the RBC is the sum of retained and estimated discards. The grey shading for year 2022 is used for stock status and RBC determination.

| Year | RBC | Retained | Discard | Status |
| :---: | :---: | :---: | :---: | :---: |
| 2017 | 2026 | 1698 | 328 | 0.87 |
| 2018 | 2010 | 1665 | 345 | 0.98 |
| 2019 | 7370 | 6914 | 456 | 1.09 |
| 2020 | 12,513 | 12,151 | 362 | 1.23 |
| 2021 | 12,341 | 12,151 | 190 | 1.41 |
| 2022 | 23,777 | 23,532 | 245 | 1.55 |
| 2023 | 21,605 | 21,391 | 214 | 1.47 |
| 2024 | 18,712 | 18,504 | 207 | 1.31 |
| 2025 | 15,848 | 15,643 | 205 | 1.14 |
| 2026 | 13,480 | 13,277 | 203 | 0.97 |
| 2027 | 11,684 | 11,482 | 201 | 0.84 |
| 2028 | 10,380 | 10,181 | 199 | 0.74 |
| 2029 | 9,458 | 9,262 | 196 | 0.66 |
| 2030 | 8,816 | 8,623 | 194 | 0.61 |
| 2031 | 8,370 | 8,178 | 191 | 0.58 |
| 2032 | 8,055 | 7,866 | 189 | 0.55 |
| 2033 | 7,827 | 7,640 | 188 | 0.54 |
| 2034 | 7,658 | 7,472 | 187 | 0.52 |
| 2035 | 7,529 | 7,343 | 186 | 0.51 |
| 2036 | 7,429 | 7,244 | 185 | 0.51 |
| 2037 | 7,351 | 7,166 | 184 | 0.50 |
| 2038 | 7,289 | 7,105 | 184 | 0.50 |
| 2039 | 7,241 | 7,058 | 184 | 0.49 |
| 2040 | 7,204 | 7,020 | 183 | 0.49 |
| 2041 | 7,174 | 6,991 | 183 | 0.49 |
| 2042 | 7,151 | 6,968 | 183 | 0.49 |
| 2043 | 7,133 | 6,950 | 183 | 0.48 |
| 2044 | 7,118 | 6,936 | 183 | 0.48 |
| 2045 | 7,107 | 6,925 | 183 | 0.48 |



Figure 6.10. The estimated time-series of relative spawning biomass and annual recruitment for the 2021 base case assessment for Blue Grenadier.


Figure 6.11. Time series showing the stock recruitment curve, recruitment deviations, recruitment deviation variance check and bias ramp for Blue Grenadier.


Figure 6.12. Kobe plot showing relative fishing mortality (y-axis) versus relative spawning biomass (x-axis).


Figure 6.13. A retrospective of assessment outputs of female spawning biomass from each stock assessment from 2001 to 2018. Note that for 2001 and 2002 only values of biomass in 1979 were available (from CastilloJordán and Tuck, 2018a).

### 6.6.2 Likelihood profiles

As stated by Punt (2018), likelihood profiles are a standard component of the toolbox of applied statisticians. They are most often used to obtain a $95 \%$ confidence interval for a parameter of interest. Many stock assessments "fix" key parameters such as $M$ and steepness based on a priori considerations. Likelihood profiles can be used to evaluate whether there is evidence in the data to support fixing a parameter at a chosen value. If the parameter is within the entire range of the $95 \%$ confidence interval, this provides no support in the data to change the fixed value. If the fixed value is outside the $95 \%$ confidence interval, it would be reasonable for a review panel to ask why the parameter was fixed and not estimated, and if the value is to be fixed, on what basis and why should what amounts to inconsistency with the data be ignored. Integrated stock assessments include multiple data sources (e.g., commonly catch-rates, length-compositions, and age-compositions) that may be in conflict, due for example to inconsistencies in sampling, but more commonly owing to incorrect assumptions (e.g., assuming that catch-rates are linearly related to abundance), i.e. modelmisspecification. Likelihood profiles can be used as a diagnostic to identify these data conflicts (Punt, 2018).

Likelihood profiles for key parameters of interest such as female natural mortality ( $M_{f}$ ), virgin spawning biomass and stock status are provided in Figure 6.14-Figure 6.16.

For Blue Grenadier, the likelihood profile for female natural mortality, $M_{f}$, is shown in Figure 6.14, with the total likelihood shown in black and components of the total likelihood from different data sources shown in a range of colours. This parameter is estimated in the model $\left(M=0.23 \mathrm{yr}^{-1}\right)$ and the likelihood profile suggests that it is reasonably well estimated, with a likely range between 0.21 and $0.26 \mathrm{yr}^{-1}$. The index and age data (suggest higher mortality) and the length data (suggest lower mortality) are in conflict. The non-spawning CPUE and to a lesser extent the egg survey data are driving the preference towards higher estimates of $M_{\mathrm{f}}$, while there is little information in the Acoustic Survey data. All length data inputs are suggesting lower estimates of $M_{\mathrm{f}}$, however, this is mostly driven by the spawning fleet onboard data. There is conflict in age data between the fleets, with the spawning fleet age data suggesting higher estimates of $M_{\mathrm{f}}$ are preferable, while the non-spawning fleet age data suggests lower estimates.

A likelihood profile for virgin spawning biomass $\left(S S B_{0}\right)$ is shown in Figure 6.15, with the total likelihood shown in black and components of the total likelihood from different data sources shown in a range of colours. This likelihood profile suggests a range of plausible values for $S S B_{0}$ ranging between around 27,000 and $52,000 \mathrm{t}$ with the most likely value at around $37,000 \mathrm{t}$. The components of the likelihood relating to the surveys suggest larger values of $S S B_{0}$ whereas the age data want lower values of $S S B_{0}$. Similarly, a likelihood profile on stock status (2020) suggests a broad range of plausible values, from approximately 0.8 to 1.7 (Figure 6.16). The index and age data suggest higher relative biomass whereas the length data suggest lower relative biomass.


Figure 6.14. The likelihood profile (top) for female natural mortality, with $95 \%$ CIs for $M_{f}$ ranging from 0.21 to 0.26 . The estimated value for $M$ is $0.23 \mathrm{yr}^{-1}$. Piner plot (bottom) for the likelihood profile showing components of the change in likelihood for index, discard, length and age in addition to the changes in the total likelihood.


Figure 6.15. The likelihood profile (top) for virgin spawning biomass, with $95 \%$ CIs ranging from $27,000 \mathrm{t}$ to $52,000 \mathrm{t}$. The estimated value is $37,000 \mathrm{t}$. Piner plot (bottom) for the likelihood profile showing components of the change in likelihood for index, discard, length and age in addition to the changes in the total likelihood.


Figure 6.16. The likelihood profile (top) for 2020 stock status, with $95 \%$ CIs ranging from 0.8 to 1.7. The estimated value is 1.25 . Piner plot (bottom) for the likelihood profile showing components of the change in likelihood for index, discard, length and age in addition to the changes in the total likelihood.

### 6.6.3 Retrospectives

A retrospective analysis was completed, starting from the most recent year of data, working backward in time and removing five successive years of data from the assessment. This analysis can highlight potential problems and instability in an assessment (Cadrin and Vaughan, 1997; Mohn, 1999). The severity of retrospective patterns can be quantified using a statistic called Mohn's rho, which is defined as the average of the relative differences between an estimate from an assessment with a truncated time series and an estimate of the same quantity from an assessment using the full time series (HurtadoFerro et al., 2015). Mohn's rho values are calculated for a range of effects, including SSB, recruitment, $F$ and stock status. As a general rule, values of Mohn's rho higher than 0.20 or lower than -0.15 are cause for concern in an assessment (Hurtado-Ferro et al., 2015). The retrospective analysis for relative and absolute spawning biomass, fit to non-spawning catch rate, and recruitment is shown in Figure 6.17, with the base case model in dark blue, and then successive years data removed back to 2015 (shown in red).

There is some evidence of over-optimistic estimation of the spawning biomass in the last year of the SSB trajectory in each case, which is also supported by Mohn's Rho being 0.26 for biomass, -0.49 for recruitment, -0.1 for $F$ and 0.26 for stock status. Of these, estimates for biomass, recruitment and stock status are higher or lower than threshold values and indicate retrospective patterns of concern, suggesting some misspecification within this assessment.


Figure 6.17. Retrospectives for relative and absolute spawning biomass, CPUE and recruitment for Blue Grenadier, with the most recent base case assessment shown (blue) and then successive years removed back to 2015 (red).

### 6.6.4 Jitter analysis

Jitter analysis is a technique used to test the optimality, robustness and stability of the maximum likelihood estimate obtained for a particular model. This involves randomly changing the starting values used for all estimated parameters and re-running the model, to test what alternative solutions may be found by the optimisation algorithm from different initial locations, which is sometimes referred to as sensitivity to initial conditions. Two diagnostics are of interest with a jitter analysis, initially a check on whether a better "optimal solution" may be found, with a higher likelihood value, and also to see how frequently the optimal solution is found. As all estimated parameters are randomly modified, or "jittered," simultaneously, this can sometimes result in a model either failing to converge or finding a local maximum in a different (suboptimal) part of the multi-dimensional parameter space. A jitter analysis was conducted with 25 replications, modifying initial values by 0.1 .

For the base case eight of the 25 jitter replicates found the same optimum solution, with a likelihood of 1922.81. The remaining 17 replicates found worse 'optimal' solutions with 16 replicated with a likelihood of 1923.24 and the last with a likelihood of 1930.00.

### 6.6.5 Sensitivities

Results of the sensitivities to the potential base case are listed in Table 6.9. The usual set of sensitivities are provided (which includes sensitivities on natural mortality, steepness, $\sigma_{R}$ and halving and doubling the weighting on length, age and index data). Relative spawning biomass varies between 1.35 and 2.12 of virgin biomass, but with most sensitivities near 1.6.

Unweighted likelihood components for the base case and differences for the sensitivities are shown in Table 6.10. This table tends to show that for most alternatives, the fit to the data is degraded by moving away from base case model values or weighting schemes.

Table 6.9. Summary of results for the base case model BC and sensitivity tests. RBC 2022-24 is the average 3year RBC. RBC 2022-26 is the average 5-year RBC. Note that only the base case is tuned.

| Model | SB0 | SB_Curr | CurrDepl | 2022 <br> RBC | RBC <br> $2022-2024$ | RBC <br> 2022-2026 | RBC <br> Long-term |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Base Case Model $\left(M_{\mathrm{f}}=0.23\right.$, |  |  |  |  |  |  |  |
| $\left.M_{\mathrm{m}}=0.24, h=0.75\right)$ | 37,445 | 57,991 | 1.55 | 23,777 | 21,365 | 18,684 | 7,100 |
| $M_{\mathrm{f}}=0.21$ | 36,245 | 48,939 | 1.35 |  |  |  |  |
| $M_{\mathrm{f}}=0.25$ | 38,442 | 65,679 | 1.71 |  |  |  |  |
| $h=0.65$ | 39,149 | 69,311 | 1.77 |  |  |  |  |
| $h=0.85$ | 38,350 | 66,991 | 1.75 |  |  |  |  |
| $\sigma_{R}=0.6$ | 34,745 | 48,002 | 1.38 |  |  |  |  |
| $\sigma_{R}=0.8$ | 42,079 | 84,083 | 2.00 |  |  |  |  |
| Double weight on Index data | 43,313 | 91,726 | 2.12 |  |  |  |  |
| Half weight on Index data | 32,439 | 44,700 | 1.38 |  |  |  |  |
| Double weight on Length data | 38,551 | 72,653 | 1.88 |  |  |  |  |
| Half weight on Length data | 39,971 | 70,952 | 1.78 |  |  |  |  |
| Double weight on Age data | 35,639 | 61,653 | 1.73 |  |  |  |  |
| Half weight on Age data | 41,872 | 69,796 | 1.67 |  |  |  |  |

Table 6.10. Summary of likelihood components for the base-case BC and sensitivity tests. Likelihood components are unweighted, and sensitivities from the BC are shown as differences from the base case. A negative value indicates a better fit, a positive value a worse fit.

| Model | TOTAL | Survey | Discard | Length <br> comp | Age <br> comp | Recruitment |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Base Case Model $\left(M_{\mathrm{f}}=0.23\right.$, | 1922.81 | -6.73 | 25.55 | 308.00 | 1505.08 | 66.61 |
| $\left.M_{\mathrm{m}}=0.24, h=0.75\right)$ | 0.97 | 0.86 | -0.01 | -1.16 | 1.53 | -0.57 |
| $M_{\mathrm{f}}=0.21$ | 0.87 | -0.53 | 0.03 | 1.61 | -0.56 | 0.59 |
| $M_{\mathrm{f}}=0.25$ | 1.44 | 0.32 | 7.91 | -9.84 | 0.27 | 2.40 |
| $h=0.65$ | -0.34 | 0.72 | 7.45 | -9.89 | 0.00 | 0.91 |
| $h=0.85$ | 21.41 | 0.74 | 2.62 | 1.14 | 4.79 | 12.35 |
| $\sigma_{R}=0.6$ | -7.65 | -3.02 | -2.10 | 9.81 | -5.15 | -6.98 |
| $\sigma_{R}=0.8$ | 8.51 | -5.54 | 2.25 | 8.83 | 0.14 | 2.95 |
| Double weight on Index data | 1.49 | 4.43 | 6.24 | -9.52 | -0.28 | 0.09 |
| Half weight on Index data | 14.04 | 0.10 | 24.20 | -41.84 | 23.55 | 4.76 |
| Double weight on Length data | 18.63 | -1.07 | -9.22 | 50.88 | -18.75 | -0.57 |
| Half weight on Length data | 12.46 | 0.80 | 2.32 | 27.02 | -29.93 | 10.35 |
| Double weight on Age data | 11.58 | -0.29 | 6.61 | -24.43 | 38.63 | -7.34 |
| Half weight on Age data |  |  |  |  |  |  |

### 6.7 Discussion

The estimated virgin female biomass is $37,445 \mathrm{t}$ (compared to $53,909 \mathrm{t}$ in 2018 and $36,815 \mathrm{t}$ in the 2013 assessments). Initial biomass is known to be sensitive in this model and often has varied between $35,000 \mathrm{t}$ and $60,000 \mathrm{t}$. The likelihood profiles reinforce that initial biomass is uncertain, as is the estimate of current stock status. However, all model sensitivities showed current relative biomass being well above the target and likely to be above initial biomass levels. There continues to be strong estimates of recent recruitment (eight years above average) which is a good sign for the fishery. As with all assessments, recent estimates of recruitment are generally less well estimated (as there are less data to inform those estimates) and so some caution should be taken with regard to the estimated recent recruitments. In addition, reducing the broad estimates of relative current biomass would be beneficial, and additional acoustic estimates of spawning biomass will likely assist in this regard. As has been observed in previous assessments of Blue Grenadier, the fit to the non-spawning fishery catch rate, especially in the early years, is poor. Further refinement of the model should consider alternative GLM models for CPUE standardisation, or potential changes to model structure to account for the poor fit. The assessment shows retrospetive patterns of concern for biomass, $F$ and stock status estimates. These results suggest that there could be some misspecification in the assessment with a time varying factor that may not be accounted for in the assessment. Further investigation of these patterns in future assessments is warranted.

Assessment outcome:
The projected 2022 spawning stock biomass will be $155 \%$ of virgin female spawning biomass (projected assuming 2020 catches in 2021), compared to $122 \%$ for 2019 in the 2018 assessment, and $94 \%$ for 2014 in the 2013 assessment.

For the base case model, the 2022 recommended biological catch (RBC) under the 20:35:48 harvest control rule is $23,777 \mathrm{t}$, with 245 t estimated discards ( $23,532 \mathrm{t}$ retained). The long-term RBC is 7,100 t , with 183 t discards.

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### 6.10 Appendix

## Length comps, retained, SpawnFleetonboard



Figure 6.18. Length composition fits: onboard spawning fleet retained.


Figure 6.19. Length composition fits: onboard non-spawning fleet retained.

Length comps, discard, NonSpawnFleetonboard


Figure 6.20. Length composition fits: onboard non-spawning fleet discard.

Length comps, retained, SpawnFleetport


Figure 6.21. Length composition fits: port spawning fleet retained.

Length comps, retained, NonSpawnFleetport


Figure 6.22. Length composition fits: port non-spawning fleet retained.


Figure 6.23. Length composition fit diagnostics from tuning. Francis data weighting method TA1.8: thinner intervals (with capped ends) show result of further adjusting sample sizes based on suggested multiplier (with $95 \%$ interval) for length data.

Pearson residuals, comparing across fleets




Figure 6.24. Residuals from the annual length compositions for base case

Conditional AAL plot, retained, SpawnFleetonboard


Figure 6.25. Fits to conditional age at length data.

Conditional AAL plot, retained, SpawnFleetonboard









Conditional AAL plot, retained, SpawnFleetonboard


Conditional AAL plot, retained, SpawnFleetonboard









Conditional AAL plot, retained, SpawnFleetonboard









Conditional AAL plot, retained, SpawnFleetonboard


Conditional AAL plot, retained, SpawnFleetonboard


Conditional AAL plot, retained, SpawnFleetonboard


Conditional AAL plot, retained, SpawnFleetonboard



Conditional AAL plot, retained, NonSpawnFleetonboard


Conditional AAL plot, retained, NonSpawnFleetonboard


Conditional AAL plot, retained, NonSpawnFleetonboard


Conditional AAL plot, retained, NonSpawnFleetonboard


Conditional AAL plot, retained, NonSpawnFleetonboard


Conditional AAL plot, retained, NonSpawnFleetonboard


Conditional AAL plot, retained, NonSpawnFleetonboard


Conditional AAL plot, retained, NonSpawnFleetonboard


Conditional AAL plot, retained, NonSpawnFleetonboard



Figure 6.26. Data weighting of conditional age at length data for the onboard non spawning and spawning fleets

Pearson residuals, retained, SpawnFleetonboard (max=24.1)


Figure 6.27. Pearson residuals of conditional age at length data.

Pearson residuals, retained, SpawnFleetonboard (max=24.1)


Pearson residuals, retained, SpawnFleetonboard (max=24.1)


Pearson residuals, retained, SpawnFleetonboard (max=24.1)


Pearson residuals, retained, SpawnFleetonboard (max=24.1)


Pearson residuals, retained, SpawnFleetonboard (max=24.1)


Pearson residuals, retained, NonSpawnFleetonboard (max=22.77)


Pearson residuals, retained, NonSpawnFleetonboard ( $\max =22.77$ )


Pearson residuals, retained, NonSpawnFleetonboard (max=22.77)


Pearson residuals, retained, NonSpawnFleetonboard (max=22.77)


Pearson residuals, retained, NonSpawnFleetonboard (max=22.77)


Pearson residuals, retained, NonSpawnFleetonboard (max=22.77)


Pearson residuals, retained, NonSpawnFleetonboard (max=22.77)


Age (yr)

